# Inference Examination

May 27, 2008 9:00 to 11:00

- 1. Answer all problems. Write only on one side of the paper.
- 2. Start each problem on a new sheet of paper with your name.
- 3. Number of points for each problem is shown in parentheses. Maximum possible is 100.

## 1. (20 points)

- (a) Define M-Estimator and Z-Estimator, and write the following three as M- and Z-Estimators: Least squares estimator, maximum likelihood estimator, sample median. (assume a random sample of n iid random variable is available)
- (b) Assume the consistency of the Z-estimator, state a theorem on asymptotic normality of Z-Estimators and outline a proof.

Mention the necessary assumptions where they are needed in the proof.

## 2. (30 points)

The following table contains the risk  $R(\theta_i, d_j)$  for different combinations of parameters  $\theta_i$ , i = 1, 2, and decision rules  $d_j$ , j = 1, ..., 4.

- (a) Draw the risk set, and mark the set of admissible rules as well as the set of lower boundary points.
- (b) Find the Bayes rule(s) and the (minimal) Bayes risk for the prior  $\pi = (4/5, 1/5)$ , where  $P_{\pi}(\theta = \theta_1) = 4/5$  and  $P_{\pi}(\theta = \theta_2) = 1/5$ .
- (c) Is this Bayes rule admissible? Why?
- (d) Find the least favorable prior distribution. Verify that it is least favorable.
- (e) Assume that there is another rule  $d_5$  with  $R(\theta_1, d_5) = 0$ ,  $R(\theta_2, d_5) = 5$ . Find the Bayes rule(s) and Bayes risk for the prior in (c).
- (f) Consider the randomized rule  $\delta$  that chooses each of  $d_2$  and  $d_5$  with probability 1/2. Under which conditions can we find a nonrandomized rule that is at least as good as  $\delta$ ?

### 3. (25 points)

Let  $(X_n)$  be a sequence of i.i.d. random variables with exponential  $(\lambda)$  distribution.

- (a) Derive the limit distribution of the minimum  $X_{(1),n}$ , after proper standardization.
- (b) Derive the limit distribution of the maximum  $X_{(n),n}$ , after proper standardization.

Hint: Proper standardizations are of the form  $(X_{(k),n} - a_n)/b_n$  for k = 1, n, and sequences  $a_n$ ,  $b_n$ , so that the limiting distribution is non-degenerate.

Repeat the above questions with  $(X_n)$  a sequence of i.i.d. random variables with Uniform[0,1] distribution.

#### 4. (25 points)

Suppose  $X_1, \dots, X_n$  are iid random variables with a CDF F(t). Assume F(t) is continuous. Denote by  $\hat{F}_n(t)$  the empirical distribution.

(a) Show that

$$\sup_{t} |\hat{F}_n(t) - F(t)|$$

converge to zero in probability as  $n \to \infty$ . i.e. for any  $\epsilon > 0$ ,

$$P(\sup_{t} |\hat{F}_n(t) - F(t)| > \epsilon)$$

goes to zero as  $n \to \infty$ .

(b) For a fixed a, we want to estimate the conditional distribution

$$P(X_1 \le t | X_1 > a) = \frac{F(t) - F(a)}{1 - F(a)}, \quad t > a.$$

Consider the estimator

$$\frac{\hat{F}_n(t) - \hat{F}_n(a)}{1 - \hat{F}_n(a)} = \hat{G}_n(t).$$

Assume F(a) < 1, show that

$$\sup_{t>a} |\hat{G}_n(t) - P(X_1 \le t | X_1 > a)|$$

converge to zero almost surely as  $n \to \infty$ .

(c) For a fixed t > a, find the asymptotic distribution of

$$\sqrt{n} \left[ \hat{G}_n(t) - P(X_1 \le t | X_1 > a) \right]$$

as  $n \to \infty$ .