

Example of using R functions related to `emplik` (empirical likelihood) package

Example 1. Find the upper and lower confidence limit of a Wilks confidence interval related to the Kaplan-Meier estimator

We can use the function `el.cen.EM` from the `emplik` package to accomplish this.

In fact `el.cen.EM` is designed to test the hypothesis (via likelihood ratio test)

$$H_0: \int g(t) dF(t) = \mu; \quad \text{vs.} \quad H_a: \int g(t) dF(t) \neq \mu$$

where $g(t)$ is a user supplied function. Depending on the $g(t)$ we supply, this hypothesis can be the test about the Kaplan-Meier mean [when $g(t) = t$], or about the survival probability at time t_0 [when $g(t) = I[t \leq t_0]$] or many others.

Let us see a real example: we use the `myeloma` data set that comes with `emplik` package

```
data(myeloma)
survtimes <- myeloma[,1]
censtatus <- myeloma[,2]
```

We shall test and construct confidence interval for $F(10)$, so define

```
myfun1 <- function(t){ as.numeric(t <= 10) }
```

Now a test based on the empirical likelihood ratio for $H_0: F(10) = 0.2$

```
el.cen.EM(fun=myfun1, x=survtimes, d=censtatus, mu=0.2)
```

among the output, you see a `Pval` (that is the P-value) of 0.3103897 so the conclusion is "not to reject the null hypothesis of $F(10)=0.2$ ".

If you want the (NPMLE) estimate of $F(10)$, look in the output for "`funMLE`"

You will see a MLE of 0.2526136 (this is just the Kaplan-Meier at $t=10$)

There is another function `el.cen.EM2` in the `emplik` package which is very similar. The function `el.cen.EM` is designed for a single parameter, and `el.cen.EM2` can handle multiple parameters. `el.cen.EM` may be a bit faster, while `el.cen.EM2` may be a bit more

robust.

For the subtle differences, please see the help pages in R.

Now we can further find the 95% confidence interval for $F(10)$. For this purpose

we need to define a function that just returns the $-2LLR$. The first input of this

function must be the value to be tested.

```
myloglikR1 <- function(mu0, x, d) {  
  myfun1 <- function(t){as.numeric(t<=10)}  
  el.cen.EM(fun=myfun1, x=x, d=d, mu=mu0)$"-2LLR"  
}
```

The real search for the lower(upper) confidence bound is done by the

```
function findL (findU)
```

```
findL(fun=myloglikR1, MLE=0.25, x=survtimes, d=censtatus)  
[1] 0.1566809230 0.0000000001 3.8399999183
```

```
findU(fun=myloglikR1, MLE=0.25, x=survtimes, d=censtatus)  
[1] 0.3684037790 0.0000000001 3.8399999664
```

There are 3 output values, the first one is the lower(upper) confidence bound,

the second/third is the final step value and the final $-2LLR$ value.

So the 95%

confidence interval for $F(10)$ is [0.156680923, 0.368403779]

If you want the 90% confidence interval instead, do this

```
findL(fun=myloglikR1, MLE=0.25, x=survtimes, d=censtatus, level=  
2.705543)
```

```
findU(fun=myloglikR1, MLE=0.25, x=survtimes, d=censtatus, level=  
2.705543)
```

Finally, the functions `findL` and `findU` is here: (not inside the `emplik` package yet)

```
findU <- function(step=0.1, initStep=0, fun, MLE=0, level=3.84, ...) {  
  value <- 0  
  Ubeta <- MLE + initStep  
  for( i in 1:8 ) {  
    while(value < level) {  
      Ubeta <- Ubeta + step  
      value <- fun(Ubeta, ... )  
    }  
    Ubeta <- Ubeta - step
```

```

    step <- step/10
    value <- fun(Ubeta, ... )
  }
return( c(Ubeta, step, value) )
}

```

```

> findL
function(step=0.1, initStep=0, fun, MLE=0, level=3.84, ...) {
#### This function try to find the 1 dim confidence interval
#### (the lower confidence limit only)
#### the fun should be the -2loglik function with 1 parameter
#### (the other parameters needs to be maxed over)
#### The initStep is provided to save some initial steps, i.e. if you
know
#### approximate where the confidence bound is, (how far it is from the
MLE)
#### then you can start search from there, instead of starting from MLE.
#### You should always use a lower bound for the initStep (a
conservative value)
#### as the search is from inside the confidence interval outwards.
value <- 0
Lbeta <- MLE - initStep
for( i in 1:8 ) {
while(value < level) {
Lbeta <- Lbeta - step
value <- fun(Lbeta, ... )
}
Lbeta <- Lbeta + step
step <- step/10
value <- fun(Lbeta, ... )
}
return( c(Lbeta, step, value) )
}

```

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In the above example, the Wald confidence interval can at least be computed easily, with the Greenwood formula to estimate the variance. The next example then is almost impossible to get a Wald confidence interval. But the Wilks confidence interval is as easy as in example one.

Example 2. Find the upper and lower confidence limit of the residual median, where the estimator is based on the Kaplan-Meier estimator.
[residual median? or median residual?]

First the definition: The residual median at a given time t_0 is the number

"theta" that solve the equation

$$\frac{1-F(t_0 + \theta)}{1-F(t_0)} = 0.5$$

The estimator of theta can be obtained by (a) replace F by the Kaplan-Meier in the above and (b) solve for theta.

Notice the variance of the estimator of theta is very complicated, and even if you get a theoretical expression for it, it is of not much use, due to the fact that it involves density function, a quantity hard to estimate.

To use the Wilks method, we shall re-cast the problem into a testing hypothesis problem.

Notice that a theta solves the above equation iff it solves

$$\int g_{\theta}(t) dF(t) = 0.5$$

where $g_{\theta}(t) = I[t \leq t_0 + \theta] - 0.5 * I[t \leq t_0]$
(simple algebraic manipulations)

and recall this hypothesis can be handled by the el.cen.EM function.

R-code:

```
library(survival)
data(cancer)
library(emplik)
cancertimes <- cancer$time
cancerstatus <- cancer$status - 1
```

This estimates the median and mean residual time at
t0=365.25 days

```
MMRtime(x=cancertimes, d=cancerstatus, age=365.25)
$MeanResidual
[1] 275.9997

$MedianResidual
[1] 258.75
```

Now the test that theta = 258

```
mygfun22 <- function(s, age, Mdage) {as.numeric(s <= (age + Mdage)) - 0.5
*as.numeric(s <= age)}
```

```
el.cen.EM(x=cancertimes, d=cancerstatus, fun=mygfun22, mu=0.5, age=
365.25, Mdage=258)$"-2LLR"
```

If we are lazy, we can just repeat the last line but change 258 to something else, until you get "-2LLR" as close to as possible (but not exceed 3.84) to find the 95% confidence interval. Or do it automatically as follows.

To find the upper and lower confidence limit by findL/findU, we need to write a function that returns the "-2LLR" and the first input variable is the value to be tested.

```
myloglikR2 <- function(mu0, x, d) {
  mygfun22 <- function(s, age=365.25, Mdage=mu0){as.numeric(s<=
(age+Mdage))-0.5*as.numeric(s<=age)}
  el.cen.EM(x=x, d=d, fun=mygfun22, mu=0.5)$"-2LLR"
}
```

Finally, we may do this (but it took too long...slow search algorithm the default search step is too small.) takes too long: findL(fun=myloglikR2, MLE=258, x=time, d=status, level=2.705543)

```
Better to do this (with bigger steps, search is a bit faster)
findL(step=5,fun=myloglikR2, MLE=258, x=time, d=status, level=2.705543)
[1] 184.75000000 0.00000005 2.50369332
findU(step=5,fun=myloglikR2, MLE=258, x=time, d=status, level=2.705543)
[1] 321.75000000 0.00000005 2.42779336
```

So, The 90% confidence interval is [184.750, 321.750].

If the step is too large, the hypothesis to be tested may be out of the range for el.cen.EM to handle and may stop computing or too slow. (here use el.cen.EM2 may be more robust)

Notice the final level reached is 2.503 or 2.427 and somewhat away from 2.7055.

THis is normal, since this is a discrete case when testing the median. (or any quantiles)

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Example 3:

The following is another example of how the code findL and findU works with parametric MLE: (the parametric regression of survreg() and getting the confidence interval for the slope in the regression)

```
> findL(fun=loglikfun, MLE=-1.255066, y=myel$dur, d=myel$status, x1=myel
$treat, x2=myel$renal)
[1] -2.7499611050  0.0000000001  3.8399999981
```

So, the lower 95% confidence bound is -2.74996

The two functions involved: findL is general (as above), loglikfun is data specific.

The MLE value can be obtained by first run a survreg, and getting the estimate, it is not critical, any value near -1.255 will work.

```
> loglikfun
function(beta, y, d, x1, x2) {
loglikmax <- survreg(Surv(y,d)~ x1 + x2)$loglik[2]
loglikH0 <- survreg(Surv(y,d)~ offset(beta*x1)+ x2)$loglik[2]
return( -2*(loglikH0 - loglikmax) )
}
```

Assumeing of course the data frame myel is already loaded there.

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Now getting the interval for both regression slopes:

```
> findU(fun=loglikfun2, MLE=-4.18, level=2.7, y=myel$dur, d=myel$status,
x1=myel$treat, x2=myel$renal)
[1] -3.1652483440  0.0000000001  2.6999999991
> findL(fun=loglikfun2, MLE=-4.18, level=2.7, y=myel$dur, d=myel$status,
x1=myel$treat, x2=myel$renal)
[1] -5.2571687790  0.0000000001  2.6999999965
>
> findU(fun=loglikfun, MLE=-1.2, level=2.7, y=myel$dur, d=myel$status,
x1=myel$treat, x2=myel$renal)
[1] -0.2800085190  0.0000000001  2.6999999965
> findL(fun=loglikfun, MLE=-1.2, level=2.7, y=myel$dur, d=myel$status,
```

```
x1=myel$treat, x2=myel$renal)
[1] -2.4572724310  0.0000000001  2.6999999991
>
```

```
> loglikfun
function(beta0, y, d, x1, x2) {
loglikmax <- survreg(Surv(y,d)~ x1 + x2)$loglik[2]
loglikH0 <- survreg(Surv(y,d)~ offset(beta0*x1)+ x2)$loglik[2]
return( -2*(loglikH0 - loglikmax) )
}
>
> loglikfun2
function(beta0, y, d, x1, x2) {
loglikmax <- survreg(Surv(y,d)~ x1 + x2)$loglik[2]
loglikH0 <- survreg(Surv(y,d)~ x1 + offset(beta0*x2))$loglik[2]
return( -2*(loglikH0 - loglikmax) )
}
>
```

=====

Example 4. Test the equality of two medians with right censored data.

The emplik package contains a function ROCnp() that can be used to test the equality of two medians from two independent samples.

```
library(survival)
times1 <- cancer$time[cancer$sex==1]
status1 <- cancer$status[cancer$sex==1] -1
times2 <- cancer$time[cancer$sex==2]
status2 <- cancer$status[cancer$sex==2] -1

ROCnp(t1=times1, d1=status1, t2=times2, d2=status2, b0=0.5, t0=0.5)
$`-2LLR`
[1] 11.02725

$cstar
[1] 312.105
```

which lead us to reject the null hypothesis of equal median (-2LLR larger than 3.84)

If we want to construct a confidence interval for the RATIO of the two medians, we need to work harder. (no existing code) But the idea is similar:

Let us test the hypothesis that $m1/m2 = 0.6$ where $m1$ is the median for male

and m_2 is the median for female. This is same as testing
 $m_1 = 0.6*m_2 = a$ for some a , or $m_1=a$, $0.6*m_2=a$; or $m_1=a$, $m_2=a/0.6$
.

We shall modify the code of `ROCnp` to accomplish this.

```
myROCnp(t1, d1, t2, d2, b0=0.5, t0=0.5)
```

```
> myROCnp
function (t1, d1, t2, d2, b0, t0)
{
  if (length(b0) != 1)
    stop("check length of b0")
  if (length(t0) != 1)
    stop("check length of t0")
  if (b0 >= 1 | b0 <= 0)
    stop("check the value of b0")
  if (t0 >= 1 | t0 <= 0)
    stop("check the value of t0")
  tempnp2 <- WKM(x = t2, d = d2)
  place2 <- sum(tempnp2$surv >= t0)
  c2 <- tempnp2$times[place2]
  tempnp1 <- WKM(x = t1, d = d1)
  place1 <- sum(tempnp1$surv > b0)
  c1 <- tempnp1$times[place1]
  if (c2 <= c1)
    c1 <- tempnp1$times[place1 + 1]
  llr <- function(const, t1, d1, t2, d2, b0, t0) {
    npllik1 <- el.cen.EM2(x = t1, d = d1, fun = function(x,
      theta) {
        as.numeric(x <= theta)
      }, mu = 1 - b0, theta = const)$"-2LLR"
    npllik2 <- el.cen.EM2(x = t2, d = d2, fun = function(x,
      theta) {
        as.numeric(x <= theta/0.6)
        ##### here is the
ratio
      }, mu = 1 - t0, theta = const)$"-2LLR"
    return(npllik1 + nppllik2)
  }
  temp <- optimize(f = llr, lower = min(c2, c1), upper = max(c2,
    c1), t1 = t1, d1 = d1, t2 = t2, d2 = d2, b0 = b0, t0 = t0)
  cstar <- temp$minimum
  val <- temp$objective
  list(`-2LLR` = val, cstar = cstar)
}
```

```
=====
```

Due to the discreteness of median, the `optimize()` function may not always find the min. An exhaust search can be done: [slower but always find the min]

```

ROCnp2 <- function (t1, d1, t2, d2, b0, t0)
{
  if (length(b0) != 1)
    stop("check length of b0")
  if (length(t0) != 1)
    stop("check length of t0")
  if (b0 >= 1 | b0 <= 0)
    stop("check the value of b0")
  if (t0 >= 1 | t0 <= 0)
    stop("check the value of t0")
  tempnp2 <- WKM(x = t2, d = d2)
  place2 <- sum(tempnp2$surv >= t0)
  c2 <- tempnp2$times[place2]
  tempnp1 <- WKM(x = t1, d = d1)
  place1 <- sum(tempnp1$surv > b0)
  c1 <- tempnp1$times[place1]
  if (c2 <= c1)
    {c1 <- tempnp1$times[place1 + 2]
     c2 <- tempnp2$times[place2 - 1] }
  llr <- function(const, t1, d1, t2, d2, b0, t0) {
    npllik1 <- el.cen.EM(x = t1, d = d1, fun = function(x,
      theta) {
        as.numeric(x <= theta)
      }, mu = 1 - b0, theta = const)$"-2LLR"
    npllik2 <- el.cen.EM(x = t2, d = d2, fun = function(x,
      theta) {
        as.numeric(x <= theta)
      }, mu = 1 - t0, theta = const)$"-2LLR"
    return(npllik1 + nppllik2)
  }
  lower <- min(c1, c2)
  upper <- max(c1, c2)
  timesALL <- c(tempnp2$times, tempnp1$times)
  midpts <- timesALL[(timesALL>lower) & (timesALL<upper)]
  midpts <- sort(midpts)
  midpts <- (midpts[-1] + midpts[-length(midpts)])/2
  k <- length(midpts)
  val2 <- rep(-9, k)
  for(i in 1:k) val2[i] <- llr(const=midpts[i], t1=t1, d1=d1, t2=t2,
d2=d2, b0=b0, t0=t0)
  val <- min(val2)
  cstar <- midpts[ which(val2==val) ]
  list(`-2LLR` = val, cstar = cstar)
}

```

=====

A similar but easier problem: test the equality of the survival probabilities at a given time t_0 ; based on two independent samples. The problem is cont. and well behaved. The minimum is easy to find.

Since the code is similar to ROCnp, I call it ROCnp3 [lack of a better

name??]

```
ROCnp3 <- function (t1, d1, t2, d2, t0)
{
  tempnp2 <- WKM(x = t2, d = d2)
  place2 <- sum(tempnp2$times <= t0)
  c2 <- tempnp2$surv[place2]
  tempnp1 <- WKM(x = t1, d = d1)
  place1 <- sum(tempnp1$times <= t0)
  c1 <- tempnp1$surv[place1]
  llr <- function(const, t1, d1, t2, d2, t0) {
    npllik1 <- el.cen.EM(x = t1, d = d1, fun = function(x) {
      as.numeric(x <= t0)
    }, mu = const)$"-2LLR"
    npllik2 <- el.cen.EM(x = t2, d = d2, fun = function(x) {
      as.numeric(x <= t0)
    }, mu = const)$"-2LLR"
    return(npllik1 + nppllik2)
  }
  lower <- min(1-c1, 1-c2)
  upper <- max(1-c1, 1-c2)
  temp <- optimize(f=llr, lower=lower, upper=upper, t1=t1, d1=d1, t2=t2,
d2=d2, t0=t0)
  cstar <- temp$minimum
  val <- temp$objective
  list(`-2LLR` = val, cstar = cstar, Prob1=1-c1, Prob2=1-c2)
}
```

The code runs like this: first I assume we took the cancer data set from package survival. take the male and female as two samples. The resulting -2LLR should be compared to $\text{chisq}(df=1)$ so we reject the H_0 of equal prob at $t_0=200$ days

```
> ROCnp3(t1=times1, d1=status1, t2=times2, d2=status2, t0=200)
$`-2LLR`
[1] 8.913334

$cstar
[1] 0.3200662

$Prob1
[1] 0.3926928

$Prob2
[1] 0.2054065
```

=====

Some thoughts when testing the ratio of two probabilities or medians.

If the actual ratio is further away from 1 then the value being tested (as in H_0), then we may search within the two prob/medians. Since we have to push the two together to make them equal or having the ratio stipulated in H_0 .

(i.e. if the two MLE's has ratio 0.5, and we are testing if the ratio is equal to 0.6 Or if the two MLE's has ratio 1.4 and we are testing H_0 ratio = 1.2)

We know where to search the common value.

When the value tested (in H_0) is further away from 1 compared to the ratio of MLE's, then we must search out side the interval formed by two MLE's. How much outside, will depends on the actual value to be tested.

If we are testing ratio equal to 2, then we should search from value a and value b, where a = larger of MLE/2, b= smaller of MLE x2. etc.

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A final note: When define median, it may be better to use a "smoothed" indicator function.

Due to discreteness nature of the Kaplan-Meier estimator. The median may be non-unique. (even after the smoothing)

Here is the indicator function and 2 smoothed version of it.

```
myfun6 <- function(x, theta) {as.numeric(x <= theta)}
```

```
myfun7 <- function(x, theta=0, eps=0.2) {
  if(eps <= 0) stop("eps must > 0")
  u <- (x-theta)/eps
  return( pmax(0, pmin(1-u, 1)) )
}
```

```
myfun5 <- function(x, theta=0, eps=0.1) {
  if(eps <= 0) stop("eps must > 0")
  u <- (x-theta)*sqrt(5)/eps
  INDE <- (u < sqrt(5)) & (u > -sqrt(5))
  u[u >= sqrt(5)] <- 0
  u[u <= -sqrt(5)] <- 1
  y <- 0.5 - (u - (u)^3/15)*3/(4*sqrt(5))
  u[ INDE ] <- y[ INDE ]
  return(u)
}
```